

Importance Sampling

PUBH 8442: Bayes Decision Theory and Data Analysis

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Importance Sampling

- ▶ Recall: Direct sampling can be used to estimate

$$\int c(\theta)p(\theta) d\theta \approx \frac{1}{N} \sum_{i=1}^N c(\theta_i)$$

by simulating draws from $p(\theta)$

- ▶ What if normalizing constant for $p(\theta)$ is unknown?
- ▶ Several “indirect” simulation methods, including *importance sampling*
- ▶ Basic idea of importance sampling:
 - ▶ Simulate θ from some other distribution g , but weigh each simulation by its relative likelihood under p .

Importance Sampling

- ▶ Consider the posterior $p(\theta | \mathbf{y}) \propto p(\mathbf{y} | \theta)p(\theta)$

- ▶ For a given function c ,

$$E(c(\theta) | \mathbf{y}) = \frac{\int c(\theta) p(\mathbf{y} | \theta) p(\theta) d\theta}{\int p(\mathbf{y} | \theta) p(\theta) d\theta}$$

Handwritten notes: $p(\theta | \vec{y})$ above the numerator, and a circle around the denominator.

- ▶ Simulate $\theta_1, \dots, \theta_N$ iid from density $g(\theta)$

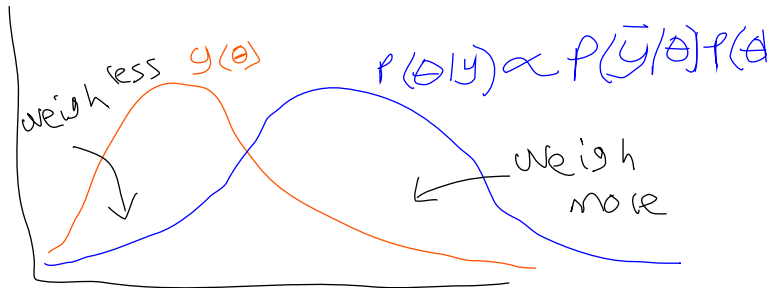
- ▶ Define the weights $w(\theta) = p(\mathbf{y} | \theta)p(\theta)/g(\theta)$:

$$E(c(\theta) | \mathbf{y}) = \frac{\int c(\theta) w(\theta) g(\theta) d\theta}{\int w(\theta) g(\theta) d\theta}$$

Handwritten notes: $p(\vec{y} | \theta) p(\theta)$ above the numerator, and a circle around $w(\theta)g(\theta)$ in the numerator.

$$\approx \frac{\frac{1}{N} \sum_{j=1}^N c(\theta_j) w(\theta_j)}{\frac{1}{N} \sum_{j=1}^N w(\theta_j)}$$

- ▶ $g(\theta)$ is called the *importance function*
- ▶ $w(\theta)$ weighs values according to their (relative) posterior probability, and probability of being simulated from g
 - ▶ Higher posterior probability \rightarrow higher weight
 - ▶ Higher probability of being simulated from $g \rightarrow$ lower weight
- ▶ Ideally, g closely resembles $p(\theta | \mathbf{y})$
 - ▶ $g = p(\theta | \mathbf{y})$ minimizes simulation variability (equal weights)
 - ▶ If $g(\theta)$ and $p(\theta | \mathbf{y})$ are very different, many weights will be close to 0 so need large N for good approximations
 - ▶ g and $p(\theta | \mathbf{y})$ should (at least) have common support.



Density approximation

- ▶ Plugging in $c(\theta) = \mathbb{1}_{\theta \in (a,b)}$ gives

$$P(\theta \in (a, b)) \approx \frac{\sum_{\theta_j \in (a,b)} w(\theta_j)}{\underbrace{\sum_{j=1}^N w(\theta_j)}}_1$$

- ▶ Can use this to approximate the posterior density with weighted histograms.

$$\text{bin height} = \frac{1}{b-a}$$


- ▶ Define

$$\widehat{c(\theta)} = \frac{\frac{1}{N} \sum_{j=1}^N c(\theta_j) w(\theta_j)}{\frac{1}{N} \sum_{j=1}^N w(\theta_j)}$$

- ▶ Define $x_j = c(\theta_j)w(\theta_j)$ and $y_j = w(\theta_j)$
- ▶ We can compute the (simulation) variability in $\widehat{c(\theta)}$ as

$$\text{Var} \left(\frac{\bar{x}}{\bar{y}} \right) \approx \frac{1}{N} \left(\frac{\hat{\sigma}_x^2}{\bar{y}^2} + \frac{\bar{x}^2 \hat{\sigma}_y^2}{\bar{y}^4} - \frac{2\bar{x} \hat{\sigma}_{xy}}{\bar{y}^3} \right)$$

where $\hat{\sigma}_x^2 = \frac{1}{N-1} \sum (x_j - \bar{x})^2$, $\hat{\sigma}_y^2 = \frac{1}{N-1} \sum (y_j - \bar{y})^2$, and

$$\hat{\sigma}_{xy} = \frac{1}{N-1} \sum (x_j - \bar{x})(y_j - \bar{y})$$

- ▶ *The expression in Carlin&Louis p.116 is incorrect*

Example: Normal-Gamma model

- ▶ Suppose $\mathbf{y} = y_1, \dots, y_n \stackrel{iid}{\sim} \text{Normal}(0, \theta)$
- ▶ Prior $\theta \sim \text{Gamma}(3, 0.5)$ ¹
- ▶ The posterior is

$$p(\theta | \mathbf{y}) \propto \theta^{(4-n)/2} \exp \left\{ -\frac{1}{2\theta} \left(\theta^2 + \sum_{i=1}^n y_i^2 \right) \right\}$$

- ▶ Not a well-known or easily integrable pdf (unlike IG prior)

¹Example credit – D. Bandyopadhyay

Example: Normal-Gamma model

- ▶ Importance sample with $g = \text{Gamma}(\alpha = cs^2, \beta = c)$
 - ▶ s^2 is the sample variance of \mathbf{y}
 - ▶ Expected value under g is s^2
 - ▶ Higher $c \rightarrow$ less variance

- ▶ Importance weight function is $w(\theta) = \frac{p(\theta) P(\vec{y}|\theta)}{q(\theta)}$

$$w(\theta) = \theta^{\frac{4-n}{2} - cs^2 + 1} \exp \left\{ -\frac{1}{2\theta} \left(\theta^2 + \sum_{i=1}^n y_i^2 \right) + c\theta \right\}$$

- ▶ Consider posterior simulation precision for different c .

Example: Normal-Gamma model

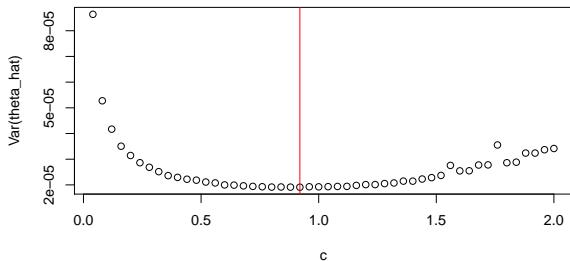
- ▶ For different values of c :
- ▶ Simulate $N = 100000$ draws from $g(\theta)$
- ▶ Compute $w(\theta)$ from each draw
- ▶ For the estimated expected value

$$\hat{\theta} = \frac{\frac{1}{N} \sum_{j=1}^N \theta_j w(\theta_j)}{\frac{1}{N} \sum_{j=1}^N w(\theta_j)}$$

compute the simulation uncertainty $\text{Var}(\hat{\theta})$ using formula.

Example: Normal-Gamma model

- Simulation variance of $\hat{\theta}$ for c for given data y_1, \dots, y_{20}



http://www.ericfrazierlock.com/More_on_Importance_Sampling_Rcode1.r

- Minimized at $c = 0.92$

Example: Normal-Gamma model

- ▶ Using draws under $c = 0.92$,

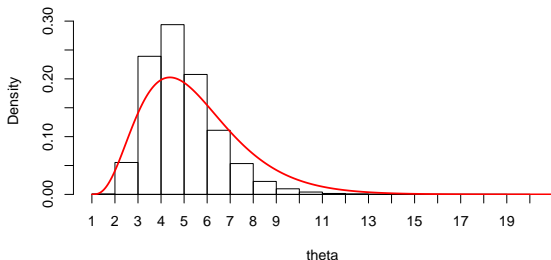
$$\hat{\theta} = 4.890$$

- ▶ We can also compute posterior probabilities, e.g.,

$$P(\theta > 5 \mid \mathbf{y}) = \frac{\sum_{\theta_j > 5} w(\theta_j)}{\sum_{j=1}^N w(\theta_j)} = 0.410$$

Example: Normal-Gamma model

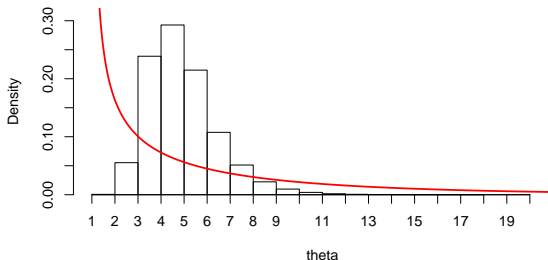
- Weighted histogram approximating posterior $p(\theta | \mathbf{y})$, for $c = 0.92$. $g(\theta)$ is shown in red



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Example: Normal-Gamma model

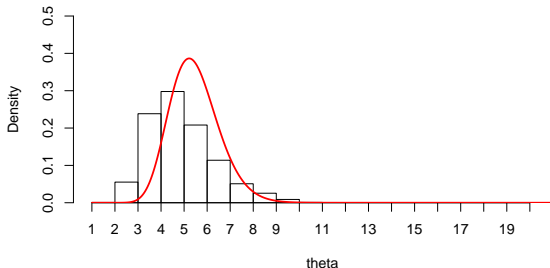
- Weighted histogram approximating posterior $p(\theta | \mathbf{y})$, for $c = 0.10$. $g(\theta)$ is shown in red



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Example: Normal-Gamma model

- Weighted histogram approximating posterior $p(\theta | \mathbf{y})$, for $c = 4$. $g(\theta)$ is shown in red



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