

## More on Linear Models

PUBH 8442: Bayes Decision Theory and Data Analysis

Eric F. Lock  
UMN Division of Biostatistics, SPH  
elock@umn.edu

03/11/2024

## Example: Body Fat

- ▶ Use iid normal prior for  $\beta'_i$ s:

$$\beta \sim \text{Normal}(0, \sigma^2 \tau^2 I)$$

- ▶  $IG(a, b)$  prior for  $\sigma^2$
- ▶ Choose  $a$  and  $b$  so that

$$E(\sigma^2) = 10, \quad \text{Var}(\sigma^2) = 100$$

$$\rightarrow a = 3, b = 20.$$

- ▶ How to choose the hyper-parameter  $\tau^2$ ?
- ▶ Could use subjective intuition, or put a prior on  $\tau^2$
- ▶ Alternatively, choose  $\tau^2$  “empirically”

$$\hat{\tau}^2 = \underset{\tau^2}{\operatorname{argmax}} p(\mathbf{y} | \tau^2)$$

- ▶ Fix  $\tau^2 = \hat{\tau}^2$  and proceed as if it were known.
- ▶ An “empirical Bayes” approach

# Normal-inverse-gamma regression marginal

- ▶ For the normal-inverse-gamma regression model with  $\beta_i$ 's iid with variance  $\tau^2$ ,

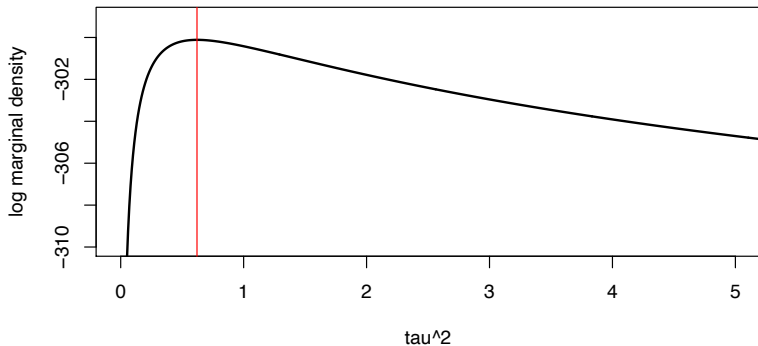
$$p(\mathbf{y} \mid \sigma^2, \tau^2) = N(\mathbf{0}, \sigma^2(I + \tau^2(\mathbf{X}\mathbf{X}^T)))$$

- ▶ The marginal for  $\mathbf{y}$  is a multivariate t-distribution with  $2a$  df:

$$p(\mathbf{y} \mid \tau^2) = MVT_{2a} \left( \mathbf{0}, \frac{b}{a}(I - \tau^2\mathbf{X}\mathbf{X}^T) \right).$$

The multivariate t density is available in R under the 'mvtnorm' package

## Example: Body Fat



[http://www.ericfrazerlock.com/More\\_on\\_Linear\\_Models\\_Rcode1.r](http://www.ericfrazerlock.com/More_on_Linear_Models_Rcode1.r)

- $p(\mathbf{y} | \tau^2)$  is maximized at  $\hat{\tau}^2 = 0.62$

## Example: Body Fat

- ▶ For hyperparameters  $\tau^2 = 0.62$ ,  $a = 3$  and  $b = 20$ , estimate

$$\tilde{\beta} = (X^T X + \frac{1}{0.62} I)^{-1} (X^T \mathbf{y})$$

- ▶ The marginal posterior for  $\sigma^2$  is

$$p(\sigma^2 | \mathbf{y}) = IG(a_n, b_n)$$

where  $a_n = 3 + \frac{100}{2}$  and

$$b_n = 20 + \frac{1}{2} [\mathbf{y}^T \mathbf{y} - \tilde{\beta}^T (X^T X + \frac{1}{0.62} I)^{-1} \tilde{\beta}]$$

- ▶ The marginal posterior for  $\beta$  is a multivariate t-distribution

$$\frac{\beta_i - \tilde{\beta}_i}{\sqrt{\frac{b_n}{a_n} ((X^T X + \frac{1}{0.62} I)^{-1})_{ii}}} \sim t_{106}.$$

## Example: Body Fat

- Estimates and 95% credible intervals for  $\beta'_i$ s:

Variable	$\tilde{\beta}_i$	95% credible interval
Age	1.193	(0.157, 2.229)
Weight	-0.465	(-4.127, 3.197)
Height	-0.380	(-1.602, 0.842)
Neck	-0.143	(-1.749, 1.463)
Chest	-0.642	(-2.999, 1.714)
Abdomen	8.704	(6.286, 11.123)
Ankle	0.042	(-1.300, 1.384)
Biceps	0.207	(-1.104, 1.519)
Wrist	-2.176	(-3.683, -0.668)

[http://www.ericfrazerlock.com/More\\_on\\_Linear\\_Models\\_Rcode1.r](http://www.ericfrazerlock.com/More_on_Linear_Models_Rcode1.r)

## Example: Body Fat (cont.)

- ▶ Recall: Predict centered  $Y = BF\%$  from 9 standardized predictors  $X$  from 100 adult males:

$$\mathbf{y} = \beta X + \epsilon$$

where  $\epsilon \sim \text{Normal}(\mathbf{0}, \sigma^2 I)$

- ▶ Considered uninformative model  $M_1$ :

$$p(\beta, \sigma^2) = \frac{1}{\sigma^2}$$

- ▶ And “shrinkage” model  $M_2$ :

$$\sigma^2 \sim IG(3, 20),$$

$$\beta \sim \text{Normal}(0, \sigma^2 \tau^2 I)$$

- ▶ Fixed  $\tau^2$  empirically:

$$\hat{\tau}^2 = \underset{\tau^2}{\operatorname{argmax}} p(\mathbf{y} | \tau^2) = 0.62.$$



## Example: Body Fat (cont.)

- ▶ Would like to compare models  $M_1$ ,  $M_2$ .
- ▶ Hard to use the Bayes factor
  - ▶ Marginal distribution not defined for  $M_1$
  - ▶ Inherent bias because  $\tau^2$  chosen to maximize marginal under  $M_2$ .
- ▶ Recall instead the log pseudo marginal likelihood:

$$LPML = \log \prod_{i=1}^n p(y_i | \mathbf{y}_{(i)})$$

where  $p(y_i | \mathbf{y}_{(i)})$  is the conditional predictive distribution for  $y_i$ .

## Example: Body Fat (cont)

- ▶ Under  $M_1$  the conditional predictive distribution is a non-central t

$$p(y_i | \mathbf{y}_{(i)}) = T_{90} \left( \mathbf{x}_i \hat{\beta}_{(i)}, s^2 (1 + \mathbf{x}_i (X_{(i)}^T X_{(i)})^{-1} \mathbf{x}_i^T) \right)$$

- ▶ Under  $M_2$  the conditional predictive distribution is a non-central t

$$p(y_i | \mathbf{y}_{(i)}) = T_{105} \left( \mathbf{x}_i \hat{\beta}_{(i)}, \frac{b_{n(i)}}{a_{n(i)}} (1 + \mathbf{x}_i (X_{(i)}^T X_{(i)} + (1/\tau^2)I)^{-1} \mathbf{x}_i^T) \right)$$

- ▶ Compute for fixed  $\tau^2 = 0.62$
- ▶ Alternatively, recompute the MLE  $\hat{\tau}^2$  based on  $\mathbf{y}_{(i)}$  for  $p(y_i | \mathbf{y}_{(i)})$  (adjusts for over-fitting).

## Example: Body Fat (cont)

- ▶ Compute
  - ▶ *LPML* for  $M_1$  is  $-288.774$
  - ▶ *LPML* for  $M_2$ , with overall  $\hat{\tau}^2$ , is  $-288.6188$
  - ▶ *LPML* for  $M_2$ , re-estimating  $\hat{\tau}^2$  for each sample, is  $-288.6155$   
[http://www.ericfrazerlock.com/More\\_on\\_Linear\\_Models\\_Rcode1.r](http://www.ericfrazerlock.com/More_on_Linear_Models_Rcode1.r)
- ▶ Suggests there (may) be a slight benefit for using the shrinkage model
- ▶ Suggests that overfitting  $\tau^2$  is not a concern.