

# Homework 6

Due by end of day Monday, 4/21

PUBH 8442: Bayes Decision Theory and Data Analysis

Include any code used to generate answers at the end of your assignment, and submit electronically via Canvas.

## 1. Mining Safety (*Modified from Exercise 3.9 in Carlin & Louis*)

Consider the count of coal mining disasters in England from 1851 to 1962, available at [http://ericfrazerlock.com/coal\\_data.txt](http://ericfrazerlock.com/coal_data.txt). We will use a hierarchical Poisson model for the number of disasters  $Y_i$  during year  $i$ , with a “changepoint” at year 1890:

$$Y_i \sim \begin{cases} \text{Poisson}(\theta) & \text{for } i = 1851, \dots, 1890 \\ \text{Poisson}(\lambda) & \text{for } i = 1891, \dots, 1962, \end{cases}$$

$$\theta \sim \text{Gamma}(1/2, b_1), \lambda \sim \text{Gamma}(1/2, b_2), \theta \text{ and } \lambda \text{ independent};$$

$$b_1 \sim \text{Gamma}(1, 1), b_2 \sim \text{Gamma}(1, 1), b_1 \text{ and } b_2 \text{ independent.}$$

- (a) Let  $\mathbf{Y} = (Y_{1851}, \dots, Y_{1962})$ . Derive the full conditional distributions for  $\theta, \lambda, b_1, b_2$ :  $p(\theta | \lambda, b_1, b_2, \mathbf{Y}), p(\lambda | \theta, b_1, b_2, \mathbf{Y}), p(b_1 | \theta, \lambda, b_2, \mathbf{Y}),$  and  $p(b_2 | \theta, \lambda, b_1, \mathbf{Y})$ .  
(Hint: they will all be Gamma distributions.)
- (b) Based on part (a), describe a Gibbs sampling algorithm to simulate from the joint posterior distribution  $p(\theta, \lambda, b_1, b_2 | \mathbf{Y})$ .
- (c) Implement your Gibbs sampler in R, and use histograms of the output to approximate the marginal posterior densities for  $\theta, \lambda$ , and  $R = \theta/\lambda$ .
- (d) Compute a 95% credible interval for  $R = \theta/\lambda$ , the ratio in mining disaster rate pre-1890 vs post-1890.

NOTE: Here the second parameter of the Gamma distribution defines the *rate*, as in class. In the Carlin & Louis book, the second parameter defines the *scale*, which is  $1/rate$ .

- a.) Let  $k$  denote the number of years until the “changepoint” (here,  $k = 40$ ). We first derive

$$\begin{aligned} P(\theta | b_1, \mathbf{y}) &\sim \text{Gamma}\left(1/2 + \sum_{i=1851}^{1890} y_i, k + b_1\right) \\ P(b_1 | \theta, \mathbf{y}) &\sim \text{Gamma}(3/2, 1 + \theta) \\ P(\lambda | b_2, \mathbf{y}) &\sim \text{Gamma}\left(1/2 + \sum_{i=1891}^{1962} y_i, n - k + b_2\right) \\ P(b_2 | \lambda, \mathbf{y}) &\sim \text{Gamma}(3/2, 1 + \lambda). \end{aligned}$$

- b.) (a) Initialize the first element of  $b_1$  and  $b_2$  to  $b_1^{(0)}$  and  $b_2^{(0)}$ .

- (b) For draws  $t=1, \dots, N$ ,

- i. Draw  $\theta^{(t)} \sim \text{Gamma}\left(1/2 + \sum_{i=1851}^{1890} y_i, k + b_1^{(t-1)}\right)$

- ii. Draw  $b_1^{(t)} \sim \text{Gamma}(3/2, 1 + \theta^{(t)})$
- iii. Draw  $\lambda^{(t)} \sim \text{Gamma}(1/2 + \sum_{i=1891}^{1962} y_i, n - k + b_2^{(t-1)})$
- iv. Draw  $b_2^{(t)} \sim \text{Gamma}(3/2, 1 + \lambda^{(t)})$

c.) Histograms for  $\theta$ ,  $\lambda$ , and  $R = \theta/\lambda$  are attached on the back page.

Following is the R program

```

# Data
coal <- read.table(url("http://ericfrazerlock.com/coal_data.txt"))
k <- 40; n <- nrow(coal)
y1 <- sum(coal[1:k, 2])
y2 <- sum(coal[(k+1):n, 2])
n1 <- k; n2 <- n-k
# Parameters
b1.init <- b2.init <- 1
# Run Gibbs
MC.n <- 10000
burn.n <- 1500
theta <- rep(NA, MC.n)
lambda <- rep(NA, MC.n)
b1 <- b2 <- rep(NA, MC.n)
pre.b1 <- b1.init; pre.b2 <- b2.init

for (i in 1:MC.n)
{
  # Draw for theta
  theta[i] <- rgamma(1, shape = 0.5+y1, rate = n1+pre.b1)
  b1[i] <- rgamma(1, shape = 1.5, rate = 1+theta[i])
  pre.b1 = b1[i]
  # Draw for lambda
  lambda[i] <- rgamma(1, shape = 0.5+y2, rate = n2+pre.b2)
  b2[i] <- rgamma(1, shape = 1.5, rate = 1+lambda[i])
  pre.b2 = b2[i]
}
R.samp <- theta/lambda

windows(width = 8)
pdf(file = "Histogram.pdf", family = "serif", width = 8)
par(mfrow=c(3,1))
hist(theta[(burn.n+1):MC.n], main = "Marginal posterior density for theta",
      xlab = expression(theta))
hist(lambda[(burn.n+1):MC.n], main = "Marginal posterior density for lambda",
      xlab = expression(lambda))
hist(R.samp[(burn.n+1):MC.n], main = "Marginal posterior density for R",
      xlab = expression(theta/lambda))

```

d.) (2.529, 4.599) computed in R by

```
round(quantile(R.samp[(burn.n+1):MC.n], probs = c(.025, .975)), 3)
```

